On the different parametrizations of the Q-exponential family distribution

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1 q-exponential family

The density (Eq. (18) of Naudt (2007)) is defined as

$$f_{\theta}(x) = c(x) \exp_{q}(-\alpha(\theta) - \theta H(x)),$$

where c, α and H are known functions. Furthermore, \exp_q is the q-deformed exponential function defined as

$$\exp_q(z) = [1 + (1-q)z]_+^{1/(1-q)}$$
 for $z \in \mathbb{R}, q \neq 1$,

where $[z]_{+} = \max(z,0)$. \exp_q is construct as the inverse of the q-deformed logarithm defined as

$$\log_q(z) = \frac{z^{1-q} - 1}{1 - q} \text{ for } z \in \mathbb{R}, q \neq 1.$$

In particular, $\forall z \in \mathbb{R}$, $\exp_q(\log_q(z)) = z$ and $\forall z \neq 0, \log_q(\exp_q(z)) = z$. Special case: for $q \to 1$, $\exp_q \to \exp$ and we get the exponential family.

Let us find the domain where 1 + (1 - q)z > 0:

• If q > 1, i.e. 1 - q < 0 then

$$1 + (1 - q)z > 0 \Leftrightarrow 1 > -(1 - q)z \Leftrightarrow \frac{-1}{1 - q} > z$$

• If q < 1, i.e. 1 - q > 0 then

$$1 + (1 - q)z > 0 \Leftrightarrow 1 > -(1 - q)z \Leftrightarrow \frac{1}{1 - q} < z$$

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2 q-Gaussian

Using

$$c(x) = 1/c_q, \ c_q = \sqrt{\frac{\pi}{1-q}} \frac{\Gamma(1+1/(1-q))}{\Gamma(3/2+1/(1-q))}, \ H(x) = x^2, \ \alpha(\theta) = \frac{\theta^{\frac{q-1}{3-q}} - 1}{q-1}, \ \theta = \sigma^{q-3},$$

we have $\alpha(\sigma) = \frac{\sigma^{q-1}-1}{q-1} = \log_{2-q}(\sigma)$. We get

$$\begin{split} f_{\sigma}(x) &= \frac{1}{c_q} \exp_q(-\log_{2-q}(\sigma) - x^2 \sigma^{q-3}) = \frac{1}{c_q} \exp_q(-\frac{\sigma^{-1+q} - 1}{-1+q} - x^2 \sigma^{q-3}) \\ &= \frac{1}{c_q} \exp_q(\frac{(1/\sigma)^{1-q} - 1}{1-q} - x^2 \sigma^{q-3}) = \frac{1}{c_q} \left[1 + (1-q) \frac{(1/\sigma)^{1-q} - 1}{1-q} - (1-q) x^2 \sigma^{q-3} \right]_+^{1/(1-q)} \\ &= \frac{1}{c_q} \left[(1/\sigma)^{1-q} - (1-q) x^2 \sigma^{q-3} \right]_+^{1/(1-q)} = \frac{1}{c_q \sigma} \left[1 - (1-q) \sigma^{1-q} x^2 \sigma^{q-3} \right]_+^{1/(1-q)} \\ &= \frac{1}{c_q \sigma} \left[1 - (1-q) x^2 / \sigma^2 \right]_+^{1/(1-q)} = \frac{1}{c_q \sigma} \exp_q(-x^2 / \sigma^2) \end{split}$$

This is different from Section 6 of Naudt (2007) where there is a typo.

3 q-Exponential

Using

$$c(x) = 1/c_q, \ c_q = \sqrt{\kappa}, \ H(x) = x, \ \alpha(\theta) = \frac{\theta^{\frac{q-1}{3-q}} - 1}{q-1}, \ \theta = \kappa^{\frac{q-3}{2}},$$

we get

$$f_{\kappa}(x) = \frac{1}{\kappa} \exp_q(-x/\kappa) = \frac{1}{\kappa} \left(1 - (1-q) \frac{x}{\kappa} \right)_+^{1/(1-q)}.$$

There exists another parametrization

$$\left\{ \begin{array}{l} \alpha+1=-1/(1-q) \\ \sigma=\alpha\kappa \end{array} \right. \Leftrightarrow \left\{ \begin{array}{l} \sigma=\alpha\kappa \\ -1/(\alpha+1)=1-q \end{array} \right. \Leftrightarrow \left\{ \begin{array}{l} \kappa=\sigma/\alpha \\ q=1+1/(\alpha+1) \end{array} \right.$$

Using the parametrization (α, σ) , we get the following density and distribution function

$$f(x) = \frac{\alpha}{\sigma} \left(1 + \frac{x}{\sigma} \right)_{+}^{-\alpha - 1}, \ F(x) = 1 - \left(1 + \frac{x}{\sigma} \right)_{+}^{-\alpha}.$$

4 Bibliography

Naudt, J. (2007), The q-exponential family in statistical physics, Journal of Physics: Conference Series 201 (2010) 012003.

Shalizi, C. (2007), Maximum Likelihood Estimation for q-Exponential (Tsallis) Distributions, http://arxiv.org/abs/math/0701854v2.